# Multivariate Statistics in Ecology and Quantitative Genetics Analyzing gene expression data or other data with more parameters than observatios

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Background Correction and Normalization for Affymetrix Microarrays

RMA

**VSN** 

Stability

Regularization

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# Affymetrix Microarrays

- Probe set for each gene
- Probe sets are spread over the chip
- ► For each probe (pm) a control probe (mm) where the

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# Robust Multi-Array Average (RMA)

### Irizarry et al., 2003

- Assume that probe expression value consists of exponential signal + normally distributed random noise.
- Each array has its own mean background level of random noise.
- After background correction, probe expression values are normalized
- ► Then pm expression values are summarized to gene expression values

### **RMA**

Model to summarize probe sets:

$$Y_{ijn} = \mu_{in} + \alpha_{jn} + \varepsilon_{ijn}$$

where n is the probe set,  $Y_{ijn}$  is are the log scaled probe expression values,  $\alpha_{jn}$  is the probe affinity effect with  $\sum_j \alpha_{jn} = 0$ , and  $\mu_{in}$  is the log expression measure for gene (probe set) n on array i.

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# Variance Stabilizing Normalization (VSN)

Huber et al., 2002

Uses error model of Rocke and Durbin (2001)

$$Y_{ijn} = \alpha_i + \beta_{ij} \cdot e^{\eta_{ijn}} + \nu_{ijn}$$

where  $\eta_{\textit{ijn}} \sim \mathcal{N}(0,1)$  and  $\nu_{\textit{ijn}} \sim \mathcal{N}(0,s_{\nu}^2)$ 

This leads to the transformation

$$h_i(y_{ijn}) = \operatorname{arsinh}(a_i + b_i y_{ijn})$$

where  $a_i$  and  $b_i$  have to be estimated from the data, and

$$\operatorname{arsinh}(x) = \log\left(x + \sqrt{x^2 + 1}\right)$$

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## Stability of gene lists

For the selection of most promising genes from gene expression data different criteria may be applied, which may lead to different list.

One way to decide which criteria are useful is to assess their stability: Do they lead to similar gene lists if part of the data is changed?

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### **Ridge Regression**

If covariables are correlated, regression coefficients  $\beta_i$  can become very large and cancel each other. A way to avoid this:

$$\widehat{\beta} = \arg\min_{\beta} \left\{ \sum_{i} \left( y_{i} - \beta_{0} - \sum_{j} x_{ij} \beta_{j} \right)^{2} + \lambda \sum_{j} \beta_{j}^{2} \right\}$$

or, in other words:

$$\widehat{\beta} = \arg\min_{\beta: \sum_{i} \beta_{i}^{2} \leq s} \left\{ \sum_{i} \left( y_{i} - \beta_{0} - \sum_{j} x_{ij} \beta_{j} \right)^{2} \right\}$$

Caution: depends on scale! Normalize the data before applying this.

### Solution for Ridge Regression:

$$\widehat{\beta} = \left( X^T X + \lambda I \right)^{-1} X^T y$$

(usual linear regression if  $\lambda = 0$ )

Alternative Motivation of Ridge Regression:  $\mathcal{N}(0, \tau^2)$ -Prior on  $\beta_j$ 

$$y_i \sim \mathcal{N}(\beta_0 + \sum_j x_{ij}\beta_j , \sigma^2)$$

Then the  $\widehat{\beta}$  are means of the posterior distribution, where  $\lambda=\sigma^2/\tau^2.$ 

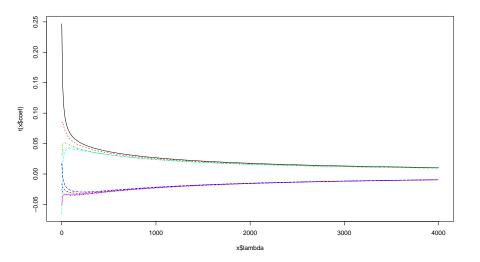
Geometric Interpretation of Ridge Regression:

All pricipal components are shrinked, the shorter they are, the more.

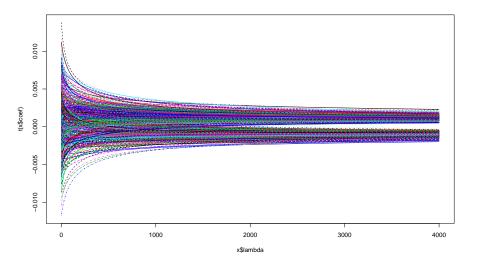
If  $d_j$  is the eigenvalue of the j-th principal component , shrink it by factor

$$rac{ extit{d}_{j}^{2}}{ extit{d}_{j}^{2}+\lambda}.$$

# Ridge regression with 10 genes



# Ridge regression with 500 genes

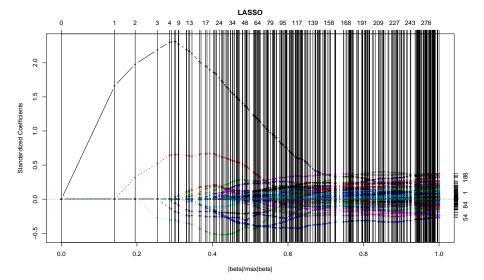


Similar method: **LASSO** (least absolute shrinkage and selector operator)

$$\widehat{\beta} = \arg\min_{\beta: \sum_{i} |\beta_{i}| \leq s} \left\{ \sum_{i} \left( y_{i} - \beta_{0} - \sum_{j} x_{ij} \beta_{j} \right)^{2} \right\}$$

the coefficients are set to 0. "kind of continous subset selection" (Hastie, Tibshirani, Friedman, 2001, *The Elements of Statistical Learning*)

# Lasso with 500 genes



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see R file